CSCI567 Machine Learning (Fall 2024)

Prof. Dani Yogatama

University of Southern California

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Outline

- **•** [Problem setup](#page-7-0)
- [K-means algorithm](#page-19-0)
- **•** [Initialization and Convergence](#page-42-0)

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Today's focus: **clustering**, an important unsupervised learning problem

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- assign each point to a specific cluster
- \bullet find the center (representative/prototype/...) of each cluster

Given: data points $\boldsymbol{x}_1, \dots, \boldsymbol{x}_N \in \mathbb{R}^{\mathsf{D}}$

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- find the cluster centers $\boldsymbol{\mu}_1,\ldots,\boldsymbol{\mu}_K \in \mathbb{R}^{\mathsf{D}}$

Many applications

- **•** recognize communities in a social network
- **•** group similar customers in market research
- **o** image segmentation
- accelerate other algorithms (e.g. NNC as in programing projects)

 \bullet ...

One example

image compression:

- each pixel is a point
- **•** perform clustering over these points
- replace each point by the center of the cluster it belongs to

Original image Large $K \longrightarrow$ Small K

Formal Objective

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Still, we can turn it into an optimization problem, e.g. through the popular "K-means" objective: find γ_{nk} and μ_k to minimize

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F(\{\gamma_{nk}\},\{\boldsymbol{\mu}_k\}) = \sum_{n=1}^N \sum_{k=1}^K \gamma_{nk} ||\boldsymbol{x}_n - \boldsymbol{\mu}_k||_2^2
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$$
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$$

The first step

$$
\min_{\{\gamma_{nk}\}} F\left(\{\gamma_{nk}\}, \{\boldsymbol{\mu}_k\}\right) = \min_{\{\gamma_{nk}\}} \sum_n \sum_k \gamma_{nk} ||\boldsymbol{x}_n - \boldsymbol{\mu}_k||_2^2
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is simply to assign each x_n to the closest μ_k , i.e.

$$
\gamma_{nk} = \mathbb{I}\left[k == \mathop{\rm argmin}\limits_{c} \|\boldsymbol{x}_n - \boldsymbol{\mu}_c\|_2^2\right]
$$

for all $k \in [K]$ and $n \in [N]$.

The second step

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$$

is simply to average the points of each cluster (hence the name)

$$
\mu_k = \frac{\sum_{n:\gamma_{nk}=1} x_n}{|\{n:\gamma_{nk}=1\}|} = \frac{\sum_{n}\gamma_{nk} x_n}{\sum_{n}\gamma_{nk}}
$$

for each $k \in [K]$.

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Step 3 Return to Step 1 if not converged

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Initialization matters for **convergence**.

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However

- it could take *exponentially many iterations* to converge
- and it *might not converge to the global minimum* of the K-means objective

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- left has K-means objective $L^2=4W^2$
- right has K-means objective W^2 , 4 times better than left!
- in fact, left is **local minimum**, and right is **global minimum**.

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• so *initialization matters a lot* for K-means

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- o or randomly assign each point to a cluster, then average: similarly fail with a constant probability
- or more sophisticated approaches: **K-means** $++$ *guarantees* to find a solution that in expectation is at most $O(\log K)$ times of the optimal

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Intuitively this *spreads out the initial centers*.

K-means++ on the same example

K-means $++$ on the same example

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So the expected K-means objective is

$$
\frac{W^2}{2(W^2+L^2)} \cdot L^2 + \left(\frac{L^2}{2(W^2+L^2)} + \frac{1}{2}\right) \cdot W^2
$$

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that is, at most 1.5 times of the optimal.

Summary for K-means

K-means is alternating minimization for the K-means objective.

The initialization matters a lot for the convergence.

 K -means $++$ uses a theoretically (and often empirically) better initialization.

Outline

2 [Gaussian mixture models](#page-78-0)

- [Motivation and Model](#page-79-0)
- [EM algorithm](#page-120-0)
- [EM applied to GMMs](#page-156-0)

Gaussian mixture models

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To solve GMM, we will introduce a powerful method for learning probabilistic model: Expectation–Maximization (EM) algorithm

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What probabilistic model generates data like this?

GMM is a natural model to explain such data

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- then draw a point according this Gaussian.

Hence the name "Gaussian mixture model".

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- \bullet μ_k and Σ_k : mean and covariance matrix of the k-th Gaussian
- \bullet N: the density function for a Gaussian

By introducing a latent variable $z \in [K]$, which indicates cluster membership, we can see p as a **marginal distribution**

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$$

x and z are both random variables drawn from the model

 \bullet x is observed

 \bullet z is unobserved/latent

An example

The conditional distributions are

$$
p(\mathbf{x} \mid z = \text{red}) = N(\mathbf{x} \mid \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1)
$$

$$
p(\mathbf{x} \mid z = \text{blue}) = N(\mathbf{x} \mid \boldsymbol{\mu}_2, \boldsymbol{\Sigma}_2)
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The marginal distribution is

 $p(x) = p(\text{red})N(x \mid \mu_1, \Sigma_1) + p(\text{blue})N(x \mid \mu_2, \Sigma_2)$ + $p(\text{green})N(\textbf{x} \mid \mu_3, \Sigma_3)$

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GMM is more explanatory than K-means

- both learn the cluster centers μ_k 's
- in addition, GMM learns cluster weight ω_k and covariance Σ_k , thus
	- we can *predict probability of seeing a new point*
	- we can generate synthetic data

How to learn these parameters?

An obvious attempt is **maximum-likelihood estimation (MLE)**: find

$$
\operatornamewithlimits{argmax}_{\boldsymbol{\theta}} \; \ln \prod_{n=1}^N p(\boldsymbol{x}_n^{\textrm{}}\,;\boldsymbol{\theta}) = \operatornamewithlimits{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^N \ln p(\boldsymbol{x}_n^{\textrm{}}\,;\boldsymbol{\theta}) \triangleq \operatornamewithlimits{argmax}_{\boldsymbol{\theta}} P(\boldsymbol{\theta})
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One solution is to still apply GD/SGD, but a much more effective approach is the Expectation–Maximization (EM) algorithm.

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Step 1 (E-Step) update the "soft assignment" (fixing parameters)

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Step 2 (M-Step) update the model parameter (fixing assignments)

$$
\omega_k = \frac{\sum_n \gamma_{nk}}{N} \qquad \mu_k = \frac{\sum_n \gamma_{nk} x_n}{\sum_n \gamma_{nk}}
$$

$$
\boldsymbol{\Sigma}_k = \frac{1}{\sum_n \gamma_{nk}} \sum_n \gamma_{nk} (\boldsymbol{x}_n - \boldsymbol{\mu}_k) (\boldsymbol{x}_n - \boldsymbol{\mu}_k)^{\mathrm{T}}
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Step 2 (M-Step) update the model parameter (fixing assignments)

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\omega_k = \frac{\sum_n \gamma_{nk}}{N} \qquad \mu_k = \frac{\sum_n \gamma_{nk} x_n}{\sum_n \gamma_{nk}}
$$

$$
\boldsymbol{\Sigma}_k = \frac{1}{\sum_n \gamma_{nk}} \sum_n \gamma_{nk} (\boldsymbol{x}_n - \boldsymbol{\mu}_k) (\boldsymbol{x}_n - \boldsymbol{\mu}_k)^{\mathrm{T}}
$$

Step 3 return to Step 1 if not converged

Step 0 Initialize $\omega_k, \mu_k, \Sigma_k$ for each $k \in [K]$

Step 1 (E-Step) update the "soft assignment" (fixing parameters)

$$
\gamma_{nk} = p(z_n = k \mid \boldsymbol{x}_n) \propto \omega_k N(\boldsymbol{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)
$$

Step 2 (M-Step) update the model parameter (fixing assignments)

$$
\omega_k = \frac{\sum_n \gamma_{nk}}{N} \qquad \mu_k = \frac{\sum_n \gamma_{nk} x_n}{\sum_n \gamma_{nk}}
$$

$$
\boldsymbol{\Sigma}_k = \frac{1}{\sum_n \gamma_{nk}} \sum_n \gamma_{nk} (\boldsymbol{x}_n - \boldsymbol{\mu}_k) (\boldsymbol{x}_n - \boldsymbol{\mu}_k)^{\mathrm{T}}
$$

Step 3 return to Step 1 if not converged

We will see how this is a special case of EM.

Generate 50 data points from a mixture of 2 Gaussians with

- $\omega_1 = 0.3, \mu_1 = -0.8, \Sigma_1 = 0.52$
- $\omega_2 = 0.7, \mu_2 = 1.2, \Sigma_2 = 0.35$

Generate 50 data points from a mixture of 2 Gaussians with

 $\omega_1 = 0.3, \mu_1 = -0.8, \Sigma_1 = 0.52$

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histogram represents the data

red curve represents the ground-truth density $p(\boldsymbol{x}) = \sum_{k=1}^K \omega_k N(\boldsymbol{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$

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•
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$$

histogram represents the data

red curve represents the ground-truth density $p(\boldsymbol{x}) = \sum_{k=1}^K \omega_k N(\boldsymbol{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$

blue curve represents the learned density for a specific round

Generate 50 data points from a mixture of 2 Gaussians with

 $\omega_1 = 0.3, \mu_1 = -0.8, \Sigma_1 = 0.52$

•
$$
\omega_2 = 0.7, \mu_2 = 1.2, \Sigma_2 = 0.35
$$

EM [demo.pdf](https://haipeng-luo.net/courses/CSCI567/2020_fall/EM_demo.pdf) shows how the blue curve moves towards red curve quickly via EM

In general EM is a heuristic to solve MLE with latent variables (not just GMM), i.e. find the maximizer of

$$
P(\bm{\theta}) = \sum_{n=1}^N \ln p(\bm{x}_n^{\texttt{}}; \bm{\theta})
$$

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$$

- θ is the parameters for a general probabilistic model
- x_n 's are observed random variables
- \bullet z_n 's are latent variables

Again, directly solving the objective is intractable.

High level idea

Keep maximizing a lower bound of P that is more manageable

Finding the lower bound of P :

$$
\ln p(\boldsymbol{x}~;\boldsymbol{\theta}) = \ln \frac{p(\boldsymbol{x}, z~;\boldsymbol{\theta})}{p(z|\boldsymbol{x}~;\boldsymbol{\theta})}
$$

(true for any z)

Finding the lower bound of P :

$$
\ln p(\boldsymbol{x} ; \boldsymbol{\theta}) = \ln \frac{p(\boldsymbol{x}, z ; \boldsymbol{\theta})}{p(z|\boldsymbol{x} ; \boldsymbol{\theta})}
$$

$$
= \mathbb{E}_{z \sim q} \left[\ln \frac{p(\boldsymbol{x}, z ; \boldsymbol{\theta})}{p(z|\boldsymbol{x} ; \boldsymbol{\theta})} \right]
$$

(true for any z)

(true for any dist. q)

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$$
\n
$$
= \mathbb{E}_{z \sim q} \left[\ln \frac{p(\boldsymbol{x}, z \; ; \boldsymbol{\theta})}{p(z|\boldsymbol{x} \; ; \boldsymbol{\theta})} \right]
$$
\n(true for any z)

\n(true for any dist. q)

$$
= \mathbb{E}_{z\sim q}\left[\ln p(\bm{x},z~;\bm{\theta})\right] - \mathbb{E}_{z\sim q}\left[\ln q(z)\right] - \mathbb{E}_{z\sim q}\left[\ln \frac{p(z|\bm{x}~;\bm{\theta})}{q(z)}\right]
$$

Finding the lower bound of P :

$$
\ln p(\boldsymbol{x} \; ; \boldsymbol{\theta}) = \ln \frac{p(\boldsymbol{x}, z \; ; \boldsymbol{\theta})}{p(z|\boldsymbol{x} \; ; \boldsymbol{\theta})}
$$
\n
$$
= \mathbb{E}_{z \sim q} \left[\ln \frac{p(\boldsymbol{x}, z \; ; \boldsymbol{\theta})}{p(z|\boldsymbol{x} \; ; \boldsymbol{\theta})} \right]
$$
\n(true for any z)

\n(true for any dist. q)

$$
= \mathbb{E}_{z \sim q} [\ln p(\boldsymbol{x}, z; \boldsymbol{\theta})] - \mathbb{E}_{z \sim q} [\ln q(z)] - \mathbb{E}_{z \sim q} \left[\ln \frac{p(z | \boldsymbol{x} ; \boldsymbol{\theta})}{q(z)} \right]
$$

$$
= \mathbb{E}_{z \sim q} [\ln p(\boldsymbol{x}, z ; \boldsymbol{\theta})] + H(q) - \mathbb{E}_{z \sim q} \left[\ln \frac{p(z | \boldsymbol{x} ; \boldsymbol{\theta})}{q(z)} \right] \quad (H \text{ is entropy})
$$

Finding the lower bound of P :

$$
\ln p(\boldsymbol{x} ; \boldsymbol{\theta}) = \ln \frac{p(\boldsymbol{x}, z ; \boldsymbol{\theta})}{p(z | \boldsymbol{x} ; \boldsymbol{\theta})}
$$
 (true for any z)

$$
= \mathbb{E}_{z \sim q} \left[\ln \frac{p(\boldsymbol{x}, z ; \boldsymbol{\theta})}{p(z | \boldsymbol{x} ; \boldsymbol{\theta})} \right]
$$
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\n
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\n(Jensen's inequality)

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\n
$$
\geq \mathbb{E}_{z \sim q} [\ln p(\boldsymbol{x}, z; \boldsymbol{\theta})] + H(q) - \ln \mathbb{E}_{z \sim q} \left[\frac{p(z | \boldsymbol{x}; \boldsymbol{\theta})}{q(z)} \right]
$$

\n(Jensen's inequality)

 $=\mathbb{E}_{z\sim q}[\ln p(\boldsymbol{x},z;\boldsymbol{\theta})]+H(q)$

Therefore, we obtain a lower bound for the log-likelihood function

$$
P(\boldsymbol{\theta}) = \sum_{n=1}^{N} \ln p(\boldsymbol{x}_n; \boldsymbol{\theta})
$$

\$\geq \sum_{n=1}^{N} (\mathbb{E}_{z_n \sim q_n} [\ln p(\boldsymbol{x}_n, z_n; \boldsymbol{\theta})] + H(q_n)) = F(\boldsymbol{\theta}, \{q_n\})\$

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This holds for any $\{q_n\}$, so how do we choose? Naturally, the one that maximizes the lower bound (i.e. the tightest lower bound)!

Equivalently, this is the same as alternatingly maximizing F over $\{q_n\}$ and θ (similar to K-means).

Fix $\boldsymbol{\theta}^{(t)}$, the solution to

$$
\operatorname*{argmax}_{q_n} \mathbb{E}_{z_n \sim q_n} \left[\ln p(\boldsymbol{x}_n, z_n \ ; \boldsymbol{\theta}^{(t)}) \right] + H(q_n)
$$

is $q_n^{(t)}$ s.t. $q_n^{(t)}(z_n) = p(z_n \mid \boldsymbol{x}_n\ ; \boldsymbol{\theta}^{(t)})$

i.e., the *posterior distribution of* z_n given \boldsymbol{x}_n and $\boldsymbol{\theta}^{(t)}$. (Verified in HW4)

Fix $\boldsymbol{\theta}^{(t)}$, the solution to

$$
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i.e., the *posterior distribution of* z_n given \boldsymbol{x}_n and $\boldsymbol{\theta}^{(t)}$. (Verified in HW4)

So at $\boldsymbol{\theta}^{(t)}$, we found the tightest lower bound $F\left(\boldsymbol{\theta},\{q^{(t)}_n\}\right)$:

Fix $\boldsymbol{\theta}^{(t)}$, the solution to

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i.e., the *posterior distribution of* z_n given \boldsymbol{x}_n and $\boldsymbol{\theta}^{(t)}$. (Verified in HW4)

So at $\boldsymbol{\theta}^{(t)}$, we found the tightest lower bound $F\left(\boldsymbol{\theta},\{q^{(t)}_n\}\right)$: $F\left(\boldsymbol{\theta},\{q_n^{(t)}\}\right) \leq P(\boldsymbol{\theta})$ for all $\boldsymbol{\theta}.$

Fix $\boldsymbol{\theta}^{(t)}$, the solution to

$$
\operatorname*{argmax}_{q_n} \mathbb{E}_{z_n \sim q_n} \left[\ln p(\bm{x}_n, z_n~;\bm{\theta}^{(t)}) \right] + H(q_n)
$$

is $q_n^{(t)}$ s.t. $q_n^{(t)}(z_n) = p(z_n \mid \boldsymbol{x}_n\ ; \boldsymbol{\theta}^{(t)}) \propto \ p(\boldsymbol{x}_n, z_n \ ; \boldsymbol{\theta}^{(t)})$

i.e., the *posterior distribution of* z_n given \boldsymbol{x}_n and $\boldsymbol{\theta}^{(t)}$. (Verified in HW4)

So at $\boldsymbol{\theta}^{(t)}$, we found the tightest lower bound $F\left(\boldsymbol{\theta},\{q^{(t)}_n\}\right)$:

\n- $$
F\left(\theta, \{q_n^{(t)}\}\right) \leq P(\theta)
$$
 for all θ .
\n- $F\left(\theta^{(t)}, \{q_n^{(t)}\}\right) = P(\theta^{(t)})$ (verify yourself by going through slide 36).
\n

Maximizing over θ

```
Fix \{q_n^{(t)}\}, maximize over \boldsymbol{\theta}:
```

$$
\operatornamewithlimits{argmax}_{\boldsymbol\theta} F\left(\boldsymbol\theta, \{q_n^{(t)}\}\right)
$$

Maximizing over θ

Fix $\{q_n^{(t)}\}$, maximize over $\boldsymbol{\theta}$:

$$
\operatorname*{argmax}_{\boldsymbol{\theta}} F\left(\boldsymbol{\theta}, \{q_n^{(t)}\}\right) \n= \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(\boldsymbol{x}_n, z_n; \boldsymbol{\theta}) \right] \quad (H(q_n^{(t)}) \text{ is independent of } \boldsymbol{\theta})
$$
Maximizing over θ

Fix $\{q_n^{(t)}\}$, maximize over $\boldsymbol{\theta}$:

$$
\operatorname*{argmax}_{\boldsymbol{\theta}} F\left(\boldsymbol{\theta}, \{q_n^{(t)}\}\right)
$$
\n
$$
= \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(\boldsymbol{x}_n, z_n; \boldsymbol{\theta}) \right] \quad (H(q_n^{(t)}) \text{ is independent of } \boldsymbol{\theta})
$$
\n
$$
\triangleq \operatorname*{argmax}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)}) \qquad (\{q_n^{(t)}\} \text{ are computed via } \boldsymbol{\theta}^{(t)})
$$

Maximizing over θ

Fix $\{q_n^{(t)}\}$, maximize over $\boldsymbol{\theta}$: argmax θ $F\left(\boldsymbol{\theta},\{q^{(t)}_{n}\}\right)$ $=$ argmax θ \sum N $\displaystyle\sum_{n=1}^{\infty}\mathbb{E}_{z_n\sim q^{(t)}_n}\left[\ln p(\bm{x}_n,z_n~;\bm{\theta})\right] ~~~(H(q^{(t)}_n)$ is independent of $\bm{\theta})$ $\triangleq \operatorname{argmax}~Q(\boldsymbol{\theta}~;\boldsymbol{\theta}^{(t)})$ θ) $(\{q_n^{(t)}\}$ are computed via $\boldsymbol{\theta}^{(t)})$

 Q is the (expected) **complete likelihood** and is usually more tractable.

Maximizing over θ

Fix $\{q_n^{(t)}\}$, maximize over $\boldsymbol{\theta}$: argmax θ $F\left(\boldsymbol{\theta},\{q^{(t)}_{n}\}\right)$ $=$ argmax θ \sum N $\displaystyle\sum_{n=1}^{\infty}\mathbb{E}_{z_n\sim q^{(t)}_n}\left[\ln p(\bm{x}_n,z_n~;\bm{\theta})\right] ~~~(H(q^{(t)}_n)$ is independent of $\bm{\theta})$ $\triangleq \operatorname{argmax}~Q(\boldsymbol{\theta}~;\boldsymbol{\theta}^{(t)})$ θ) $(\{q_n^{(t)}\}$ are computed via $\boldsymbol{\theta}^{(t)})$

 Q is the (expected) **complete likelihood** and is usually more tractable.

versus the incomplete likelihood: $P(\bm{\theta}) = \sum_{n=1}^{N} \ln p(\bm{x}_n~;\bm{\theta})$

 ${\bf Step\ 0}$ Initialize $\boldsymbol{\theta}^{(1)},\ t=1$

 ${\bf Step\ 0}$ Initialize $\boldsymbol{\theta}^{(1)},\ t=1$

Step 1 (E-Step) update the posterior of latent variables

$$
q_n^{(t)}(\cdot) = p(\cdot \mid \boldsymbol{x}_n \ ; \boldsymbol{\theta}^{(t)})
$$

 ${\bf Step\ 0}$ Initialize $\boldsymbol{\theta}^{(1)},\ t=1$

Step 1 (E-Step) update the posterior of latent variables

$$
q_n^{(t)}(\cdot) = p(\cdot \mid \boldsymbol{x}_n \; ; \boldsymbol{\theta}^{(t)})
$$

and obtain Expectation of complete likelihood

$$
Q(\boldsymbol{\theta}:\boldsymbol{\theta}^{(t)})=\sum_{n=1}^{N}\mathbb{E}_{z_n\sim q_n^{(t)}}\left[\ln p(\boldsymbol{x}_n, z_n~;\boldsymbol{\theta})\right]
$$

 ${\bf Step\ 0}$ Initialize $\boldsymbol{\theta}^{(1)},\ t=1$

Step 1 (E-Step) update the posterior of latent variables

$$
q_n^{(t)}(\cdot) = p(\cdot \mid \boldsymbol{x}_n \; ; \boldsymbol{\theta}^{(t)})
$$

and obtain **Expectation** of complete likelihood

$$
Q(\boldsymbol{\theta}:\boldsymbol{\theta}^{(t)})=\sum_{n=1}^{N}\mathbb{E}_{z_n\sim q_n^{(t)}}\left[\ln p(\boldsymbol{x}_n, z_n~;\boldsymbol{\theta})\right]
$$

Step 2 (M-Step) update the model parameter via Maximization

$$
\boldsymbol{\theta}^{(t+1)} \leftarrow \operatornamewithlimits{argmax}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}~; \boldsymbol{\theta}^{(t)})
$$

Step 3 $t \leftarrow t + 1$ and return to Step 1 if not converged

$$
P(\boldsymbol{\theta}^{(t+1)}) \ge F\left(\boldsymbol{\theta}^{(t+1)}~;\{q^{(t)}_n\}\right)
$$

$$
P(\boldsymbol{\theta}^{(t+1)}) \ge F\left(\boldsymbol{\theta}^{(t+1)}; \{q_n^{(t)}\}\right) \\
\ge F\left(\boldsymbol{\theta}^{(t)}; \{q_n^{(t)}\}\right)
$$

$$
P(\boldsymbol{\theta}^{(t+1)}) \ge F\left(\boldsymbol{\theta}^{(t+1)}; \{q_n^{(t)}\}\right) \\
\ge F\left(\boldsymbol{\theta}^{(t)}; \{q_n^{(t)}\}\right) \\
= P(\boldsymbol{\theta}^{(t)})
$$

 $P(\bm{\theta})$ is non-concave, but $Q(\bm{\theta};\bm{\theta}^{(t)})$ often is concave and easy to maximize.

$$
P(\boldsymbol{\theta}^{(t+1)}) \ge F\left(\boldsymbol{\theta}^{(t+1)}; \{q_n^{(t)}\}\right) \\
\ge F\left(\boldsymbol{\theta}^{(t)}; \{q_n^{(t)}\}\right) \\
= P(\boldsymbol{\theta}^{(t)})
$$

So EM always increases the objective value and will converge to some local maximum (similar to K-means).

E-Step:

$$
q_n^{(t)}(z_n = k) = p\left(z_n = k \mid \mathbf{x}_n ; \boldsymbol{\theta}^{(t)}\right)
$$

$$
\propto p\left(\mathbf{x}_n, z_n = k ; \boldsymbol{\theta}^{(t)}\right)
$$

E-Step:

$$
q_n^{(t)}(z_n = k) = p\left(z_n = k \mid \mathbf{x}_n; \boldsymbol{\theta}^{(t)}\right)
$$

$$
\propto p\left(\mathbf{x}_n, z_n = k; \boldsymbol{\theta}^{(t)}\right)
$$

$$
= p\left(z_n = k; \boldsymbol{\theta}^{(t)}\right) p(\mathbf{x}_n \mid z_n = k; \boldsymbol{\theta}^{(t)})
$$

E-Step:

$$
q_n^{(t)}(z_n = k) = p\left(z_n = k \mid \boldsymbol{x}_n; \boldsymbol{\theta}^{(t)}\right)
$$

$$
\propto p\left(\boldsymbol{x}_n, z_n = k \mid \boldsymbol{\theta}^{(t)}\right)
$$

$$
= p\left(z_n = k \mid \boldsymbol{\theta}^{(t)}\right) p(\boldsymbol{x}_n \mid z_n = k \mid \boldsymbol{\theta}^{(t)}\right)
$$

$$
= \omega_k^{(t)} N\left(\boldsymbol{x}_n \mid \boldsymbol{\mu}_k^{(t)}, \boldsymbol{\Sigma}_k^{(t)}\right)
$$

E-Step:

$$
q_n^{(t)}(z_n = k) = p\left(z_n = k \mid \mathbf{x}_n; \boldsymbol{\theta}^{(t)}\right)
$$

$$
\propto p\left(\mathbf{x}_n, z_n = k; \boldsymbol{\theta}^{(t)}\right)
$$

$$
= p\left(z_n = k; \boldsymbol{\theta}^{(t)}\right) p(\mathbf{x}_n \mid z_n = k; \boldsymbol{\theta}^{(t)})
$$

$$
= \omega_k^{(t)} N\left(\mathbf{x}_n \mid \boldsymbol{\mu}_k^{(t)}, \boldsymbol{\Sigma}_k^{(t)}\right)
$$

This computes the "soft assignment" $\gamma_{nk}=q^{(t)}_n(z_n=k)$, i.e. conditional probability of x_n belonging to cluster k.

M-Step:

$$
\operatorname*{argmax}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) = \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(\boldsymbol{x}_n, z_n; \boldsymbol{\theta}) \right]
$$

M-Step:

$$
\operatorname*{argmax}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) = \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(\boldsymbol{x}_n, z_n; \boldsymbol{\theta}) \right]
$$

$$
= \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(z_n; \boldsymbol{\theta}) + \ln p(\boldsymbol{x}_n | z_n; \boldsymbol{\theta}) \right]
$$

M-Step:

$$
\operatorname*{argmax}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) = \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(\boldsymbol{x}_n, z_n; \boldsymbol{\theta}) \right]
$$

$$
= \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(z_n; \boldsymbol{\theta}) + \ln p(\boldsymbol{x}_n | z_n; \boldsymbol{\theta}) \right]
$$

$$
= \operatorname*{argmax}_{\{\omega_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k\}} \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma_{nk} \left(\ln \omega_k + \ln N(\boldsymbol{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right)
$$

M-Step:

$$
\operatorname*{argmax}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}) = \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(\boldsymbol{x}_n, z_n; \boldsymbol{\theta}) \right]
$$

$$
= \operatorname*{argmax}_{\boldsymbol{\theta}} \sum_{n=1}^{N} \mathbb{E}_{z_n \sim q_n^{(t)}} \left[\ln p(z_n; \boldsymbol{\theta}) + \ln p(\boldsymbol{x}_n | z_n; \boldsymbol{\theta}) \right]
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$$
= \operatorname*{argmax}_{\{\omega_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k\}} \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma_{nk} \left(\ln \omega_k + \ln N(\boldsymbol{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right)
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To find $\omega_1, \ldots, \omega_K$, solve

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To find each μ_k , Σ_k , solve

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$$

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Solutions to previous two problems are very natural, for each k

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i.e. (weighted) fraction of examples belonging to cluster k

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You will verify some of these in HW4.

EM for learning GMMs:

Step 0 Initialize ω_k , μ_k , Σ_k for each $k \in [K]$

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Step 2 (M-Step) update the model parameter (fixing assignments)

$$
\omega_k = \frac{\sum_n \gamma_{nk}}{N} \qquad \boldsymbol{\mu}_k = \frac{\sum_n \gamma_{nk} \boldsymbol{x}_n}{\sum_n \gamma_{nk}}
$$

$$
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T

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Step 3 return to Step 1 if not converged

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- when $\sigma \rightarrow 0$, EM becomes K-means

GMM is a soft version of K-means and it provides a probabilistic interpretation of the data, which means we can predict and generate data after learning.